Échantillonnage Monte-Carlo multi-niveaux pour la quantification d'incertitudes, l'analyse de sensibilité et l'assimilation de données

Application à la modélisation de l'écoulement des fleuves

Matthias De Lozzo

Post-doctorant au CERFACS (Toulouse) matthias.delozzo.fr - matthias@delozzo.fr

Vendredi 15 septembre 2017 Séminaire de l'unité de MIA de Toulouse [De Rocquigny et al., 2008]

Step C : Propagation of uncertainty sources



Figure: Uncertainty study and management in an industrial approach. Source: Bertrand looss, ENBIS-EMSE 2009 Conference, Saint-Étienne, July 2009.

- ► Travaux de recherche au CERFACS¹, Toulouse
- Avec Paul Mycek² ainsi que Sophie Ricci^{1,2}, Mélanie Rochoux^{1,2}, Pamphile Roy² et Nicole Goutal³ ¹CECI - CNRS ²CERFACS ³LHSV - EDF R&D

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 - Modélisation <u>sous incertitudes</u> de la hauteur et du débit d'un fleuve (ex. : Garonne, estuaire de la Gironde, etc.)

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 Méthodologie :
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 - 3. Réduction d'incertitudes (ex. : criblage, mise-à-jour, etc.)
 - 4. Assimilation de données (ex. : filtre de Kalman d'ensemble)

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- Deux types de solution :
 - Modèles de substitution
 - Méthodes d'échantillonnage

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Context: hydraulic modelling in presence of uncertainties

Mean estimation by Monte-Carlo sampling

Mean estimation by multilevel Monte-Carlo (MLMC) sampling

Extension of this MLMC estimator to various statistics

Extension of these MLMC tools to data assimilation

Extension of these MLMC tools to sensitivity analysis

Application: modelling the hydraulic state of the Garonne

Conclusion, perspectives and references

Miscellaneous

Application: the Garonne river between Tonneins and La Réole

Framework: hydraulic network over the Garonne river

- length: 647 km,
- drained area: 55.000 km²,
- ▶ a Garonne flood in 2009 at Marmande.



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Application: focus on the section Tonneins - La Réole

- Form Tonneins (s = 13 km) to La Réole (s = 62 km),
- observation station at Marmande (s = 13 km),
- length: 50 km,
- mean slope: 3.3 m.km⁻¹,
- mean width: 250 m,
- ▶ bank-full discharge \approx mean annual discharge = 1.000 m³.s⁻¹.

Discharge and water level, solutions of the shallow water equations

Navier-Stokes equations (NSE)

modeling the free surface flow dynamics with the hydraulic state:

 $\{Q(s,t): \text{ discharge } [m^3.s^{-1}]; h(s,t): \text{ water level } [m]\}$

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- ▶ height ≫ width,
- small bathymetry variations,
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- ▶ height ≫ width,
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- hydrostatic pressure.
- \Rightarrow Hyperbolic system of partial differential equations representing:
 - subcritical flows (flow velocity < wave velocity),
 - supercritical flows (flow velocity > wave velocity),
 - hydraulic jumps (abrupt slowing down \Rightarrow rather abrupt rise).

The software MASCARET in the field of free-surface flow



You are here: Home + Introduction + Welcome to TELEMAC-MASCARET

Welcome to TELEMAC-MASCARET

TELEMAC-MASCARET is an integrated suite of solvers for use in the field of free-surface flow. Having been used in the context of many studies throughout the world, it has become one of the major standards in its field.

TELEMAC-MASCARET is managed by a consortium of core organisations: Artelia (formerly Sogreah, France), BundesAnstalt fur Wasserbau (BAW, Germany), Centre d'Etudes et d'Expertise sur les Risques, l'Environnement, la Mobilité et l'Aménagement (CEREMA, France), Daresbury Laboratory (United Kingdom), Electricité de France RAB (EDF, France), and HR Wallingford (United Kingdom).

TELEMAC-MASCARET is used by most partners for dimensioning and impact studies, where safety is prevailing and, for this reason, reliability, validation and a worldwide recognition of our tools are of utmost importance. As a consequence and to improve the access to TELEMAC-MASCARET for the whole community of consultants and researchers, the choice of open source has been made. Anyone can thus take advantage of TELEMAC-MASCARET and assess its performances, and will find necessary resources on this website. However the quality of assistance, maintenance and hotline support are also very important to professional users, and a special effort has been made to offer alternatively a broad range of fee-paging services.

Latest News

2017 User Conference -Registration now open ! 14 August 2017

Registration for the upcoming XXIVth TELEMAC-MASCARET Users Conference is open ! You can register for the conference under www.tuc2017.tugraz.at where you will find also the preliminary program and t [...]

07 August 2017 v7p2r2 is available for download

31 July 2017 Telemac Summer School 2017

Modelling water levels and flow rates for a river using MASCARET

Simulator: MASCARET software (using our Python API)

- simulates the one-dimensional shallow water equations,
- returns the hydraulic state (Q, h),
- over a discrete hydraulic network $\{s_{in}, s_{in} + \Delta s, \dots, s_{out}\}$,
- at different times $\{t_1, t_1 + \Delta t, \ldots, t_T\}$.

Applications: dam-break wave simulation, reservoir flushing and flooding, low flow, flood spreading, ...

MASCARET inputs

- bathymetry and roughness coefficients,
- upstream/downstream boundary conditions,
- lateral inflows,
- initial conditions.

Water level uncertainty quantification by Monte-Carlo methods



Uncertain parameters

- Strickler coefficient K_s,
- ▶ upstream flow Q_{in}.

Quantities of interest

- water level at Marmande,
- even at different sites.

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- statistics (e.g. mean or variance),
- sensitivity indices (e.g. Sobol' indices),
- excess probabilities,
- covariance matrices for data assimilation,

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Bottleneck

Variance of Monte-Carlo estimators *vs.* Simulator cost

Implementation facility vs. computational cost

The physical phenomenon

f

The considered simulator

$$\begin{array}{ll} : & \mathcal{X} \to \mathbb{R} & f_L : & \mathcal{X} \to \mathbb{R} \\ & \mathbf{x} \mapsto y = f(\mathbf{x}) & \mathbf{x} \mapsto y_L = f_L(\mathbf{x}) \end{array}$$

where *L* denotes an accuracy level assumed satisfying.

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Uncertainty framework

Input parameters $\mathbf{x} = (x_1, \dots, x_d)$ are uncertain:

▶ **x** is an instance of the r.v. $\mathbf{X} = (X_1, \dots, X_d) \sim \mathcal{L}(X_1, \dots, X_d)$

Thus, the outputs $y := f(\mathbf{x})$ and $y_L := f_L(\mathbf{x})$ are uncertain:

• y is an instance of the r.v. $Y = f(\mathbf{X}) \sim \mathcal{L}(Y) = ?$

► y_L is an instance of the r.v. $Y_L = f_L(\mathbf{X}) \sim \mathcal{L}(Y_L) = ?$

Implementation facility vs. computational cost

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MC estimation of mean, covariance, variance, ... For any application $g : \mathcal{X} \to \mathbb{R}$ and any sample size N:

1. Build a *N*-sample $(\mathbf{X}^{(i)}, g(\mathbf{X}^{(i)}))_{1 \le i \le N}$.

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$$\mathsf{Covariance:} \ \hat{c}_{N}\left[g(\mathbf{X}), g(\tilde{\mathbf{X}})\right] = \frac{1}{N-1} \sum_{i=1}^{N} \left(g\left(\mathbf{X}^{(i)}\right) - \hat{m}_{N}\right) \left(g\left(\tilde{\mathbf{X}}^{(i)}\right) - \hat{m}_{N}\right)$$

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$$\mathsf{Variance:} \ \hat{\mathsf{s}}_{N}^{2}\left[g(\mathsf{X})\right] = \frac{1}{N-1}\sum_{i=1}^{N}\left(g\left(\mathsf{X}^{(i)}\right) - \hat{m}_{N}\right)^{2}$$

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 $\Rightarrow \underline{\text{Accuracy}} \sim \mathcal{O}\left(\textit{N}^{-1/2}\right) \textit{ vs. } \underline{\text{ computational cost}} \sim \mathcal{O}\left(\textit{m,h,d,w}\right).$

e.g. MC expectation using codes with increasing accuracy and cost

Let $\{f_\ell\}_{\ell \ge 0}$ be a sequence of simulators with:

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$$Y_{\ell} \equiv f_{\ell}(\mathbf{X})$$
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Expressing the expectation as a telescoping sum

$$\mathbb{E}[Y_L] = \mathbb{E}[Y_0] + \sum_{\ell=1}^L \mathbb{E}[Y_\ell] - \mathbb{E}[Y_{\ell-1}]$$

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Expressing the expectation as a telescoping sum

$$\begin{split} \mathbb{E}[Y_{L}] &= \mathbb{E}[Y_{0}] + \sum_{\ell=1}^{L} \mathbb{E}[Y_{\ell}] - \mathbb{E}[Y_{\ell-1}] \\ \mathbb{E}_{L}^{\mathrm{ML}}[Y] &= \mathbb{E}_{0}[Y_{0}] + \sum_{\ell=1}^{L} \mathbb{E}_{\ell}[Y_{\ell}] - \mathbb{E}_{\ell}[Y_{\ell-1}] \end{split} \text{ where }$$

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Cost and variance of the MLMC estimator (Giles, 2008)

$$C \equiv \sum_{\ell=0}^{L} N_{\ell} C_{\ell} \qquad \qquad V \equiv \sum_{\ell=0}^{L} V_{\ell} / N_{\ell}$$

• C_{ℓ} = computational cost of $Y_{\ell} - Y_{\ell-1}$.

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$$V_{\ell}$$
 = variance of $Y_{\ell} - Y_{\ell-1}$.

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 \forall accuracy ε s.t. $\mathbb{E}\left[\left(E_{L}^{\mathsf{ML}}[Y] - \mathbb{E}[Y]\right)^{2}\right] < \varepsilon^{2}$, there are $N_{0}, N_{1}, \ldots, N_{\ell}$ satisfying the optimal cost:

$$C = \varepsilon^{-2} \left(\sum_{\ell=0}^{L} \sqrt{V_{\ell} C_{\ell}} \right)^2.$$

⇒ When $V_{\ell}C_{\ell}$ increases (resp. decreases), *C* is reduced by factor V_L/V_0 (resp. C_0/C_L) compared to standard MC with level *L*.

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Theorem (Giles, 2008)

Let $\{M_\ell\}_{\ell=0}^\infty \in \mathbb{N}_+$ with an exponential increase s.t. $M_\ell/M_{\ell-1} \ge a$ for some fixed a.
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Let $\{M_{\ell}\}_{\ell=0}^{\infty} \in \mathbb{N}_+$ with an exponential increase s.t. $M_{\ell}/M_{\ell-1} \ge a$ for some fixed a. Let α, β, γ be positive constants s.t.:

(i) $|\mathbb{E}[Y_{\ell} - Y_{\ell-1}]| \lesssim M_{\ell}^{-\alpha}$ (ii) $V_{\ell} \lesssim M_{\ell}^{-\beta}$ (iii) $C_{\ell} \lesssim M_{\ell}^{\gamma}$.

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Then, $\forall \varepsilon > 0$, $\exists L$, $\exists \{N_{\ell}\}_{0 \le \ell \le L}$ s.t. $\mathbb{E}\left[\left(E_{L}^{ML} - \mathbb{E}[Y]\right)^{2}\right] < \varepsilon^{2}$, the overall computational cost verifies:

$$C \lesssim \varepsilon^{-\frac{\gamma}{\alpha}} + \varepsilon^{-2} \left(\mathbb{1}_{\beta > \gamma} + |\log(\varepsilon)|^2 \, \mathbb{1}_{\beta = \gamma} + \varepsilon^{-\frac{\gamma - \beta}{2}} \mathbb{1}_{\beta < \gamma} \right)$$

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 \Rightarrow Up to a logarithmic term, the computational cost is reduced by a factor of the order of $\varepsilon^{\frac{\min(2\alpha,\beta,\gamma)}{\alpha}}$ vs. standard MC with level *L*.

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⇒ Up to a logarithmic term, the computational cost is reduced by a factor of the order of $\varepsilon^{\frac{\min(2\alpha,\beta,\gamma)}{\alpha}}$ vs. standard MC with level *L*. *Remark:* M_{ℓ} represents the "complexity" of the software at level ℓ . For example, we can choose $M_{\ell} = 2^{\ell}$ when the size of the mesh differs by a factor of 2 between levels $\ell - 1$ and ℓ .

- 1. Initialization:
 - L_{\min} , L_{\max} = minimal, maximal number of levels

 - L := L_{min} = current number of levels
 N₀^[1], N₁^[1], ..., N₁^[1] = small initial sample sizes

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Sequential selection of N_{ℓ} and L (Giles, 2015)

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2.5 If
$$N_{\ell, \text{opt}} \gg N_{\ell}$$
,

Back to Step 2. with it := it + 1

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- 2. For each iteration *it* and each level *I*:
 - 2.1 Simulate the $N_{\ell}^{[it]}$ -sample of $(Y_{\ell-1}, Y_{\ell})$.
 - 2.2 Compute empirical mean and variance of $Y_{\ell} Y_{\ell-1}$.
 - 2.3 Estimate α and β by linear regression
 - 2.4 Estimate the optimal number of simulations $N_{\ell,opt}$
 - 2.5 If $N_{\ell,\text{opt}} \gg N_{\ell}$,

Back to Step 2. with it := it + 1

Else if accuracy not reached and $L < L_{max}$

Add a finer level L := L + 1

Initialize the variance and the cost $Y_L - Y_{L-1}$

Update $N_{0.\text{opt}}, \ldots, N_{L.\text{opt}}$

Back to Step 2. with it := it + 1

to more general quantities (Bierig and Chernov, 2015)

MLMC estimator built from unbiased MC estimators Let:

• θ be the quantity to be estimated

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Then, we may defined the corresponding MLMC estimator:

$$\hat{\theta}_L^{\rm ML} = \sum_{\ell=0}^L \hat{\theta}_{\ell,\ell} - \hat{\theta}_{\ell-1,\ell}, \quad \hat{\theta}_{-1,0} \equiv 0,$$

with:

•
$$\mathbb{E}\left[\hat{\theta}_{L}^{\mathrm{ML}} - \theta_{L}\right] = 0,$$

• $\mathbb{E}\left[\hat{\theta}_{L}^{\mathrm{ML}} - \theta\right] = \theta_{L} - \theta \xrightarrow{L \to \infty} 0$

MLMC estimators of covariance and variance

For uncertainty quantification and data assimilation purposes, we are particularly interested in estimating variances and covariances.

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Unbiased covariance MLMC estimator Let Y and Z be two r.v. Then, we have:

$$C_{L}^{\mathrm{ML}}[Y, Z] = \sum_{\ell=0}^{L} C_{\ell}[Y_{\ell}, Z_{\ell}] - C_{\ell}[Y_{\ell-1}, Z_{\ell-1}]$$

where $C_{\ell}[Y_{\ell'}, Z_{\ell'}] = \frac{N_{\ell}}{N_{\ell} - 1} E_{\ell} \left[(Y_{\ell'} - E_{\ell}[Y_{\ell'}]) (Z_{\ell'} - E_{\ell}[Z_{\ell'}]) \right].$

MLMC estimators of covariance and variance

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Special case: the unbiased variance MLMC estimator

$$V_L^{\mathrm{ML}}[Y] \equiv C_L^{\mathrm{ML}}[Y,Y] = \sum_{\ell=0}^L V_\ell[Y_\ell] - V_\ell[Y_{\ell-1}]$$

where $V_{\ell}[Y_{\ell'}] \equiv C_{\ell}[Y_{\ell'}, Y_{\ell'}].$

Mean square error decomposition and bias/variance trade-off

Mean square error (MSE) of a MLMC estimator

$$\mathsf{MSE}\left(\hat{ heta}_L^{\mathrm{ML}}, heta
ight) \ \equiv \mathbb{E}\left[\left(\hat{ heta}_L^{\mathrm{ML}}- heta
ight)^2
ight]$$

Mean square error decomposition and bias/variance trade-off

Mean square error (MSE) of a MLMC estimator

$$\begin{split} \mathsf{MSE}\left(\hat{\theta}_{L}^{\mathrm{ML}},\theta\right) & \equiv \mathbb{E}\left[\left(\hat{\theta}_{L}^{\mathrm{ML}}-\theta\right)^{2}\right] \\ & = \mathbb{V}\left[\hat{\theta}_{L}^{\mathrm{ML}}\right] + \mathsf{Bias}\left(\hat{\theta}_{L}^{\mathrm{ML}},\theta\right)^{2} \end{split}$$

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Objective

Controlling the MSE of the MLMC estimator $\hat{\theta}_L^{\text{ML}}$ using a sequence of number of evaluations $\{N_\ell\}_{0 \le \ell \le L}$ balancing:

- its cost (= sampling error)
- its accuracy (= discretization error).

Bounding the variance of a MLMC estimator²

$$\begin{array}{c|c} \hat{\theta}_{L}^{\mathrm{ML}} = \dots & \mathbb{V}\left[\hat{\theta}_{L}^{\mathrm{ML}}\right] \leq \dots \\ \hline E_{L}^{\mathrm{ML}}[Y] & \sum_{\ell \leq L} \frac{1}{N_{\ell}} \mathbb{V}[\Delta_{\ell}^{Y}] \text{ (equality)} \\ \hline V_{L}^{\mathrm{ML}}[Y] & \sum_{\ell \leq L} \frac{1}{N_{\ell} - 1} \sqrt{\mathbb{M}^{4}[\Delta_{\ell}^{Y}] \mathbb{M}^{4}[\Sigma_{\ell}^{Y}]} \end{array}$$

where:

$$\begin{array}{l} \blacktriangleright \ \Delta_{\ell}^{Y} \equiv Y_{\ell} - Y_{\ell-1}, \\ \blacktriangleright \ \Sigma_{\ell}^{Y} \equiv Y_{\ell} + Y_{\ell-1}, \\ \blacktriangleright \ \mathbb{M}^{4} \text{ denotes the fourth central moment.} \end{array}$$

²For $E_L^{ML}[Y]$, see Giles, 2008. For $V_L^{ML}[Y]$, see Bierig and Chernov, 2015.

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Extending the results of Bierig and Chernov, 2015

Theorem Let Y and Z be r.v. Let Y_{ℓ} and Z_{ℓ} be their approximations at level ℓ .

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$$\begin{array}{c} (i) \left| \mathbb{C}[Y_{\ell}, Z_{\ell}] - \mathbb{C}[Y, Z] \right| \lesssim M_{\ell}^{-\alpha}, \\ (ii) \sqrt{\mathbb{M}^{4}[\Delta_{\ell}^{Y}]} + \sqrt{\mathbb{M}^{4}[\Delta_{\ell}^{Z}]} \lesssim M_{\ell}^{-\beta}, \ (iii) \ \mathcal{C}_{\ell} \lesssim M_{\ell}^{\gamma}, \end{array}$$

and assume that $\mathbb{M}^4[Y_\ell]$ and $\mathbb{M}^4[Z_\ell]$ are uniformly bounded.

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and assume that $\mathbb{M}^{4}[Y_{\ell}]$ and $\mathbb{M}^{4}[Z_{\ell}]$ are uniformly bounded. Then for any $0 < \varepsilon < e^{-1}$, there exist an integer L and a sequence of positive integers $(N_{\ell})_{\ell \geq 0}$ s.t. $\mathbb{E}\left[(C_{L}^{\mathrm{ML}}[Y, Z] - \mathbb{C}[Y, Z])^{2}\right] < \varepsilon^{2}$ and:

$$\mathsf{Cost}_{\varepsilon}\big(C_{\mathsf{L}}^{\mathrm{ML}}[Y,Z]\big) \lesssim \varepsilon^{-\frac{\gamma}{\alpha}} + \varepsilon^{-2} \left(\mathbbm{1}_{\beta > \gamma} + \mathsf{log}(\varepsilon)^2 \mathbbm{1}_{\beta = \gamma} + \varepsilon^{-\frac{\gamma - \beta}{2}} \mathbbm{1}_{\beta < \gamma}\right).$$

Numerical experiments for MLMC covariance Study case: a stochastic differential equation - Presentation

An ordinary differential equation with a random coefficient

$$\left\{ egin{aligned} &rac{\mathrm{d} u(t,\omega)}{\mathrm{d} t} = \mathsf{a}(\omega) u(t,\omega), \quad t\in(0,1], \quad \mathsf{a}\sim\mathcal{N}(\mu,\sigma^2) \ &u(0,\cdot) = U_0, \end{aligned}
ight.$$

whose solution $u(t, \cdot) = U_0 e^{at} \sim \ln \mathcal{N}(\mu t + \ln U_0, \sigma^2 t^2)$ is a log-normal process.

Statistics with analytical formulation $\mathbb{E}[u(t,\cdot)], \mathbb{V}[u(t,\cdot)], \mathbb{C}[u(t_1,\cdot), u(t_2,\cdot)], \mathbb{C}[a, u(t,\cdot)], \mathbb{M}^4[u(t,\cdot)]$ and $\mathbb{M}^4[a]$.

Definition of the different levels $M_{\ell} = 16 \times 2^{\ell}$ esquispaced points at level ℓ

Numerical experiments for MLMC covariance

Study case: a stochastic differential equation - Results





to data assimilation

Ensemble Kalman filtering

- MLMC covariance matrices based:
 - either on a direct application of the MLMC covariance, using a worst-case strategy and sparse matrices
 - either on an extension of the theorem dedicated to the covariance, using a MSE definition dedicated to covariance matrices.
- ▶ MLMC Ensemble Kalman filtering (see Hoel et al, 2016)

to sensitivity analysis

Sobol' index $S(X_i) = \mathbb{V}[\mathbb{E}[Y|X_i]]/\mathbb{V}[Y]$

1. $S(X_i) \propto \mathbb{V}[\mathbb{E}[Y|X_i]] \leftarrow \text{trick for MLMC} + \text{dimensioned value}$

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- 1. $S(X_i) \propto \mathbb{V}[\mathbb{E}[Y|X_i]] \leftarrow \text{trick for MLMC} + \text{dimensioned value}$
- 2. Rewrite $\mathbb{V}[\mathbb{E}[Y|X_i]] = \mathbb{C}[Y, Y^{\{i\}}]$ with $Y^{\{i\}} = f(X_{\sim i}, X'_i)$ where X'_i is an independent copy of X_i .

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- 3. Apply MLMC techniques for the covariance.

Extension of Janon, 2013 Let $\left\{S_{i,\Psi_{\ell}(N)}^{(\ell)}\right\}_{0 \le \ell \le L}$ be a sequence of pick-and-freeze estimators of the Sobol' index $S(X_i)$ where $\Psi_{\ell} : \mathbb{N}^* \to \mathbb{N}^*$ and $\Psi_0(N) > \Psi_1(N) > \ldots > \Psi_L(N) \equiv N$. Then, we define the multi-fidelity estimator:

to sensitivity analysis

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$$S_{i,N}^{\mathsf{ML}} = \sum_{\ell=0}^{L} S_{i,\Psi_{\ell}(N)}^{(\ell)} - S_{i,\Psi_{\ell}(N)}^{(\ell-1)}$$
Extending the MLMC techniques

to sensitivity analysis

Theorem (Concentration inequality; extension of Janon, 2013)) If $\forall \ell \in \{0, 1, ..., L\}$, $\lim_{N \to \infty} \Psi_{\ell}(N) = \infty$, Then, for any sequence of accuracy levels $\{\alpha_{\ell}\}_{0 \leq \ell \leq L}$ in]0,1[, we have:

$$\lim_{N\to\infty} \mathbb{P}\left[\left|S_{i,N}^{ML} - S_i\right| \leq \sum_{\ell=0}^{L} q(\alpha_\ell) \frac{\sigma_\ell}{\sqrt{\Psi_\ell(N)}}\right] \geq 1 - \sum_{\ell=0}^{L} \alpha_\ell$$

with $q(a) = \Phi^{-1}(1 - a/2)$

$$\left\{ \begin{array}{l} \sigma_0 = \frac{\mathbb{V}[A_0 - B_0/2]}{(\mathbb{V}[Y_0])^2} \\ \sigma_\ell = \sigma_{\ell-1} + \frac{\mathbb{V}[A_\ell - B_\ell/2]}{(\mathbb{V}[Y_\ell])^2} - \frac{2\operatorname{Cov}[A_\ell, A_{\ell-1}] - \left(\operatorname{Cov}(A_\ell, B_{\ell-1}) + \operatorname{Cov}(B_\ell, A_{\ell-1})\right) + \operatorname{Cov}(B_\ell, B_{\ell-1})/2}{\mathbb{V}[Y_\ell]\mathbb{V}[Y_\ell-1]}, \ \ell > 0 \end{array} \right.$$

with $A_\ell = (Y_\ell - \mathbb{E}[Y_\ell])(Y_\ell^{\{i\}} - \mathbb{E}[Y_\ell])$ and $B_\ell = (Y_\ell - \mathbb{E}[Y_\ell])^2 + (Y_\ell^{\{i\}} - \mathbb{E}[Y_\ell])^2$.

Application of MLMC in hydraulic modeling

Coming back to the Garonne river



Output of interest:

Water level at Marmande h

Uncertain inputs:

- Strickler coefficient K_s $K_s \sim \mathcal{N}(35, 3.6)$
- Upstream flow Q $Q \sim \mathcal{N}(1500, 102)$

MLMC parameters:

- Levels:
 - $\ell \in \{0,1,\ldots,L_{\text{max}}\}$
- Finest level: $L_{max} = 6$
- Number of nodes by level:
 - $N_{x,0} = 25$ nodes
 - $N_{x,\ell} = 2^{\ell} \times N_0$ nodes

CPU time(Y_ℓ) ~ $\mathcal{O}(2^{0.98\ell})$:

- 0. 0.2 s 4. 2.3 s
- 1. 0.3 s 5. 4.5 s
- 2. 0.6 s 6. 8.9 s
- 3. 1.1 s

Results concerning the mean of the water level



$$\begin{split} |\mathsf{Mean}\left(Y_{\ell} - Y_{\ell-1}\right)| &\sim \mathcal{O}\left(2^{-\alpha \ell}\right) \qquad |\mathsf{Variance}\left(Y_{\ell} - Y_{\ell-1}\right)| \sim \mathcal{O}\left(2^{-\beta \ell}\right) \\ \alpha &= 1.21 \qquad \qquad \beta = 1.90 \end{split}$$

Results concerning the mean of the water level



Accuracy: $\varepsilon = 5.000e - 03$

- Number of runs:
 - Level 0: 14738 runs
 - Level 1: 624 runs
 - Level 2: 268 runs
 - Level 3: 86 runs
 - Level 4: 38 runs
 - Level 5: 34 runs
 - Level 6: 13 runs
- Comparison at $\ell = 6$:
 - MLMC pprox 317 runs
 - MC \approx 11482 runs
- Savings: 36.24

Results concerning the variance of the water level



Accuracy: $\varepsilon = 1.000e - 02$

- Number of runs:
 - Level 0: 3030 runs
 - Level 1: 257 runs
 - Level 2: 146 runs
 - Level 3: 42 runs
 - Level 4: 15 runs
- Comparison at $\ell = 4$:
 - MLMC \approx 294 runs
 - MC \approx 11482 runs
- **Savings: 4.39**

Results concerning a Sobol' index



Accuracy: $\varepsilon = 1.000e - 03$

- ► Number of runs:
 - Level 0: 341489 runs
 - Level 1: 46837 runs
 - Level 2: 15202 runs
 - Level 3: 9915 runs
 - Level 4: 4273 runs
 - Level 5: 1347 runs
 - Level 6: 677 runs
- Comparison at $\ell = 6$:
 - MLMC \approx 11407 runs
 - MC \approx 97651 runs
- Savings: 8.56

Conclusion

- General framework for MLMC approach
- Extension to the covariance case:
 - Theorem
 - Adaptative algorithm
- Practical extension to the target "cost" (adaptative algorithm)
- Successfully applied to an analytical test case
 - Expectation, variance, covariance
 - Comparison of the computed and theoretical bounds
- Successfully applied to an hydraulic problem
 - Expectation, variance, covariance, numerator of Sobol' indices
 - Important reduction of the computation cost
- Straightforward extension of the multi-fidelity estimator for Sobol' indices

Perspectives

- Publication: work in progress :-)
- Extension of the MLMC covariance estimator to the case of covariance matrices
 - Data assimilation
- Extension of the MLMC pick-and-freeze Sobol' index to the case of a set of MLMC pick-and-freeze Sobol' indices
 - Sensitivity analysis
- Application of the MLMC methods to an hydraulic problem based on an unstructured mesh
- MLMC strategy for the construction of multifidelity surrogate models (e.g. cokriging)
- MLMC strategy for excess probabilities

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Ali Baba's cave: https://people.maths.ox.ac.uk/gilesm/mlmc_community.html

matthias.delozzo.fr



Qui suis-je?

Matthias DE LOZZO

30 ans, originaire de Toulouse Actuellement à Toulouse, après une expérience de deux ans en Provence

in Mon profil LinkedIn - 🏦 Mon profil ResearchGate

Ingénieur-Docteur en Modélisation Statistique

> Formation : INSA de Toulouse (ingénieur depuis 2010 - docteur depuis 2013)

> Depuis octobre 2016 : post-doctorant au CERFACS (Centre Européen de Recherche et de Formation Avancée en Calcul Scientifique)

> Avril 2014 - Mars 2016 : post-doctorant au CEA (Commissariat à l'énergie atomique et aux énergies alternatives) de Cadarache

> Qualification (MCF) - Section 26 (Mathématiques appliquées et applications des mathématiques)

> Expérience professionnelle : CNRM - Météo France, ONERA, Epsilon-Ingénierie, CEA puis CERFACS.

Attrait pour un large champ d'application des statistiques : environnement, énergie, économie, social, industrie, santé, text mining, big data ...



Métier

Compétences :

Apprentissage statistique

- Exploration de données
- Machine learning & Data mining
- Modèles de régression et de décision
- Modèles mixtes, durée de vie
- Plans d'expériences et enrichissement
- Tests statistiques
- ...

Thématiques de recherche :

- Gestion des incertitudes
- Analyse de sensibilité
- Modèles de substitution
- Métamodèles par processus gaussien

Informatique

Métier : LaTeX Matlab OpenTURNS R

Programmation : C + C++ + Fortran + Python

Web : HTML . PHP . CSS . SQL

Système : Windows + Linux

Langues

Français : langue maternelle

Anglais : professionnel

Espagnol : intermédiaire

Centres d'intérêt

Associatif ◆ Data analysis et Data visualization ◆ Environnement et Société ◆ Montagne ◆ Webmastering

matthias.delozzo.fr



2012 Comparison of Polynomial Chaos and Gaussian Process surrogates for uncertainty quantification and correlation estimation of spatially distributed open-channel steady flows, P. T. Roy, N. El Moçayd, S. Ricci, J.-C. Jouhaud, N. Goutal, M. De Lozzo, M. Rochoux, source in a structure for the search and Risk Assessment.

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- 2017 Dependence and variance-based measures for sensitivity analysis with multidimensional variables, M. De Lozzo, A.
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- 2015 Advanced surrogate model and sensitivity analysis methods for SFR accident assessment, A. Marrel, N. Marle, M. De Lozzo, ele Reliability Engineering & System Safety, 138, p. 232–241.
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Jounaux français

- 2015 Substitution de modèle et approche multifidélité en expérimentation numérique, M. De Lozzo, Journal de la Société
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Actes de conférences

- 2017 Méthodes de Monte-Carlo multi-niveaux pour la quantification d'incertitudes et l'assimilation de données Application
- I a la modélisation fluviale, M. De Lozzo, P. Mycek, S. Ricci, M. Rochoux, P. T. Roy, N. Goutal, 49èmes Journées de Statistique, SFDS, Avignon, du 29 mai au 2 juin 2017.
- 2017 Uncertainty quantification for the Gironde estuary hydrodynamics with Telemac 2D, V. Laborle, N. Goutal, S. Ricci, M. De Lozzo, P. Sergent, SymHydro 2017.
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- 2015 Uncertainty propagation for reliability analysis: application to sodium fast reactor safety studies, A. Marrel, N. Seiler Marie, M. De Lozzo, UNCECOMP, Hersonissos (Grèce), du 25 au 27 mai 2015.
- 2012 Learning of Spatio-temporal Dynamics in Thermal Engineering, M. De Lozzo, P. Klotz, B. Laurent, Engineering Applications of
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MERCI pour votre attention :-)

Matthias De Lozzo

Plus d'informations sur matthias.delozzo.fr

Assimilation de données par filtre de Kalman

Modèle d'observation $\mathbf{Y}^{o} = \mathbf{G}\mathbf{X}^{b} + \boldsymbol{\epsilon}$ où $\mathbf{X}^{b} \sim \mathcal{N} [\mathbf{X}^{t}, \mathbf{B}], \ \boldsymbol{\varepsilon} \perp \mathbf{X}^{b}$ et $\boldsymbol{\varepsilon} \sim \mathcal{N} [\mathbf{0}, \boldsymbol{R}]$ Vecteur gaussien de l'observation et de l'analyse $\begin{pmatrix} \mathbf{X}^{a} \\ \mathbf{Y}^{o} \end{pmatrix} \sim \mathcal{N} \begin{bmatrix} \begin{pmatrix} \mathbf{x}^{b} \\ \mathbf{G}\mathbf{x}^{b} \end{pmatrix}, \begin{pmatrix} \mathbf{B} & \mathbf{B}\mathbf{G}^{T} \\ \mathbf{G}\mathbf{B} & \mathbf{G}\mathbf{B}\mathbf{G}^{T} + \mathbf{R} \end{pmatrix} \end{bmatrix}$

Estimation de l'analyse (accompagnée d'une mesure d'erreur)

$$\mathbf{x}^{a} = \mathbb{E}[\mathbf{X}^{a} | \mathbf{Y}^{o} = \mathbf{y}^{o}]$$
$$= \mathbf{x}^{b} + \mathbf{B}\mathbf{G}^{T} \left(\mathbf{G}\mathbf{B}\mathbf{G}^{T} + \mathbf{R}\right)^{-1} \left(\mathbf{y}^{o} - \mathbf{G}\mathbf{x}^{b}\right)$$